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# LG–Holling type II diseased predator ecosystem with Lévy noise and white noise

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## Abstract

This paper considers an LG–Holling type II diseased predator ecosystem with Lévy noise and white noise. In this prey–predator system, assuming that the predator population is infected by disease and divided into two classes: susceptible predator and infected predator, we show that the system has a unique global positive solution. We investigate the persistence in the mean and extinction for the population, obtain threshold conditions of extinction, persistence in the mean by Itô’s formula and a comparison theorem for the stochastic system. In addition, we discuss the uniform boundedness of  $p$ th moment with  $p > 0$  and reveal the stochastically ultimate boundedness of the system. Finally, some numerical simulations are introduced to illustrate our analytical findings.

**Keywords:** Leslie–Gower; Prey–predator system; Lévy noise; Persistence in the mean; Ultimate boundedness

## 1 Introduction

In ecology, prey–predator systems are always the important part, and the most classic model is Lotka–Volterra model. Out of the ordinary, Leslie introduced the following model on the premise that the carrying capacity of predator was proportional to the amount of prey [1, 2]:

$$\begin{cases} \frac{dN_1}{dt} = (r_1 - a_1N_1 - b_2N_2)N_1, \\ \frac{dN_2}{dt} = (r_2 - b_2\frac{N_2}{N_1})N_2, \end{cases} \quad (1)$$

here  $N_1$  and  $N_2$  represented the densities of prey and predator at time  $t$ , respectively.  $r_1$  and  $r_2$  were the intrinsic growth rates for prey and predator populations.  $b_2$  was conversion rate. Leslie assumed that the predator population still grew logistically with carrying capacity  $\frac{r_2N_1}{b_2}$ .

Leslie assumed that the growth of predator population still satisfied the logistic rule, but the carrying capacity, which was distinguished from the constant in the common model, was determined by the density of prey population. Recently, there have been many experts and researchers studying or improving a Leslie type prey–predator model. Time delay, impulse, food chain, disease, environment pollution, reaction diffusion and so on, have been introduced into the model, see [3–8].

In [9], Sarwardi introduced a contagious disease among the predators in a predator-prey ecosystem, modeled via the following equations with nonnegative parameters:

$$\begin{cases} \frac{dx}{dt} = a_1x - b_1x^2 - \frac{c_1xy}{x+k_1} - \frac{pc_1xz}{x+k_1}, \\ \frac{dy}{dt} = a_2y - \frac{c_2y(y+z)}{x+k_2} - \lambda yz, \\ \frac{dz}{dt} = \lambda yz + a_3z - \frac{c_3z(y+z)}{x+k_2}. \end{cases} \tag{2}$$

The system incorporates a Holling type II functional response and a modified version of the Leslie–Gower one. In system (2), the following assumptions are made:

- (A1)  $x(t)$  represents the density of prey population at time  $t$ , and it is assumed that the disease spreads only among the predator,  $y(t)$  and  $z(t)$  denote the densities of susceptible predator and infected predator at time  $t$ , the total predator population is  $n(t) = y(t) + z(t)$ . The disease transmission rate is  $\lambda$ .
- (A2) In the absence of predators, the prey population grows logistically with intrinsic growth rate  $a_1 > 0$  and carrying capacity  $a_1b_1^{-1}$ .  $c_1$  and  $pc_1$  are the predation rates for susceptible and infected predator. Because of disease infection, the predation ability for the infected predator is a bit less than that for the susceptible predator; therefore,  $0 < p < 1$ .
- (A3) Here  $a_2, a_3$  are the per capita growth rates of each predator population. The parameters  $k_1, k_2$  respectively represent the half saturation constants of prey and predator populations. In addition, the authors introduce intraspecific competition among the predators' sound and infected subpopulations with parameters  $c_2$  and  $c_3$ , for which  $c_2 > c_3$ .

The authors studied the stability of equilibrium point of system (2) and persistence.

Population growth in the natural world is inevitably affected by the environmental perturbations. In the past decades, white noise was introduced into the model to simulate the stochastic perturbation of environment. However, sudden environmental perturbation, such as earthquakes, hurricanes, the flood, etc., cannot be modeled by white noise, which is often used to describe those stable and continuous stochastic interferences. Bao et al. in [10, 11] firstly proposed that these phenomena could be described by Lévy noise. At present, there are many research papers considering Lévy noise [6, 12–17]. In this paper, we use white noise and Lévy noise to simulate the random change of environment. The following system with jumps is considered:

$$\begin{cases} dx(t) = [a_1x(t) - b_1x^2(t) - \frac{c_1x(t)y(t)}{x(t)+k_1} - \frac{\tilde{p}c_1x(t)z(t)}{x(t)+k_1}] dt \\ \quad + \sigma_1x(t) dB_1(t) + \int_{\mathbb{Y}} \gamma_1(u)x(t^-)\tilde{N}(dt, du), \\ dy(t) = [a_2y(t) - \frac{c_2y(t)(y(t)+z(t))}{x(t)+k_2} - \beta y(t)z(t)] dt \\ \quad + \sigma_2y(t) dB_2(t) + \int_{\mathbb{Y}} \gamma_2(u)y(t^-)\tilde{N}(dt, du), \\ dz(t) = [\beta y(t)z(t) + a_3z(t) - \frac{c_3z(t)(y(t)+z(t))}{x(t)+k_2}] dt \\ \quad + \sigma_3z(t) dB_3(t) + \int_{\mathbb{Y}} \gamma_3(u)z(t^-)\tilde{N}(dt, du), \end{cases} \tag{3}$$

with the initial value  $(x_0, y_0, z_0) \in R_+^3$ . For convenience, we use  $\tilde{p}$  and  $\beta$  instead of  $p$  and  $\lambda$  in system (3), respectively.  $B_1(t), B_2(t)$  and  $B_3(t)$  are mutually independent Brownian motions defined on a complete probability space  $(\Omega, \mathcal{F}, \mathcal{P})$  with a filtration  $\{\mathcal{F}_t\}_{t \in R_+}$  satisfying the usual conditions.  $\sigma_i^2$  ( $i = 1, 2, 3$ ) denotes the intensity of white noise. In system (3),  $x(t^-), y(t^-)$  and  $z(t^-)$  are the left limits of  $x(t), y(t)$  and  $z(t)$ , respectively.  $N$  is a Poisson

counting measure with characteristic measure  $\lambda$  on a measurable subset  $\mathbb{Y}$  of  $(0, +\infty)$  with  $\lambda(\mathbb{Y}) < \infty$ ,  $\tilde{N}(dt, du) = N(dt, du) - \lambda(du) dt$ ,  $\gamma_i : \mathbb{Y} \times \Omega \rightarrow \mathbb{R}$  is bounded and continuous with respect to  $\lambda$ , and is  $\mathcal{B}(\mathbb{Y}) \times \mathcal{F}_t$ -measurable,  $i = 1, 2, 3$ .

For simplicity, we introduce the following notations:

$$w_i = a_i - \frac{\sigma_i^2}{2} - \int_{\mathbb{Y}} [\gamma_i(u) - \ln(1 + \gamma_i(u))] \lambda(du), \quad i = 1, 2, 3.$$

$$R_+^3 = \{(\eta_1, \eta_2, \eta_3)^T \in R^3 \mid \eta_i > 0, i = 1, 2, 3\}.$$

**Assumption 1** There exists a constant  $c$  satisfying

$$\int_{\mathbb{Y}} [\ln(1 + \gamma_i(u))]^2 \lambda(du) < c, \quad i = 1, 2, 3.$$

Throughout this paper, we assume that Assumption 1 is true.

**2 The existence of global positive solution**

**Theorem 1** For any initial value  $(x_0, y_0, z_0) \in R_+^3$ , system (3) has a unique global solution  $(x(t), y(t), z(t)) \in R_+^3$  for  $t \geq 0$ , and the solution will remain in  $R_+^3$  almost surely.

*Proof* At first, let us consider the following system:

$$\begin{cases} du(t) = [w_1 - b_1 \exp(u(t)) - \frac{c_1 \exp(v_1(t))}{\exp(u(t))+k_1} - \frac{\tilde{p}c_1 \exp(v_2(t))}{\exp(u(t))+k_1}] dt \\ \quad + \sigma_1 dB_1(t) + \int_{\mathbb{Y}} \ln(1 + \gamma_1(u)) \tilde{N}(dt, du), \\ dv_1(t) = [w_2 - \frac{c_2(\exp(v_1(t))+\exp(v_2(t)))}{\exp(u(t))+k_2} - \beta \exp(v_2(t))] dt \\ \quad + \sigma_2 dB_2(t) + \int_{\mathbb{Y}} \ln(1 + \gamma_2(u)) \tilde{N}(dt, du), \\ dv_2(t) = [w_3 + \beta \exp(v_1(t)) - \frac{c_3(\exp(v_1(t))+\exp(v_2(t)))}{\exp(u(t))+k_2}] dt \\ \quad + \sigma_3 dB_3(t) + \int_{\mathbb{Y}} \ln(1 + \gamma_3(u)) \tilde{N}(dt, du), \end{cases} \tag{4}$$

on  $t \geq 0$  with the initial data  $(\ln x_0, \ln y_0, \ln z_0)$ . Clearly, the coefficients of system (4) satisfy the local Lipschitz condition, then there is a unique local solution on  $[0, \tau_e)$ , where  $\tau_e$  is the explosion time. Therefore, by Itô's formula,  $(\exp(u(t)), \exp(v_1(t)), \exp(v_2(t)))$  is the unique positive local solution of system (3) with the initial value  $(x_0, y_0, z_0) \in R_+^3$ . Whereafter, we will show that  $\tau_e = \infty$ , namely the solution of system (3) is globally existent.

Consider the stochastic differential equations:

$$\begin{cases} d\Phi_1(t) = \Phi_1(t)[a_1 - b_1 \Phi_1(t)] dt + \sigma_1 \Phi_1(t) dB_1(t) + \int_{\mathbb{Y}} \gamma_1(u) \Phi_1(t^-) \tilde{N}(dt, du), \\ d\Phi_2(t) = \Phi_2(t)[a_2 - \frac{c_2 \Phi_2(t)}{\Phi_1(t)+k_2}] dt + \sigma_2 \Phi_2(t) dB_2(t) + \int_{\mathbb{Y}} \gamma_2(u) \Phi_2(t^-) \tilde{N}(dt, du), \\ d\Phi_3(t) = \Phi_3(t)[\beta \Phi_2(t) + a_3 - \frac{c_3 \Phi_3(t)}{\Phi_1(t)+k_2}] dt \\ \quad + \sigma_3 \Phi_3(t) dB_3(t) + \int_{\mathbb{Y}} \gamma_3(u) \Phi_3(t^-) \tilde{N}(dt, du), \end{cases} \tag{5}$$

and

$$\begin{cases} d\Psi_1(t) = \Psi_1(t)[a_1 - b_1 \Psi_1(t) - \frac{c_1 \Phi_2(t)}{k_1} - \frac{\tilde{p}c_1 \Phi_3(t)}{k_1}] dt \\ \quad + \sigma_1 \Psi_1(t) dB_1(t) + \int_{\mathbb{Y}} \gamma_1(u) \Psi_1(t^-) \tilde{N}(dt, du), \\ d\Psi_2(t) = \Psi_2(t)[a_2 - \frac{c_2 \Psi_2(t)}{k_2} - \frac{c_2 \Phi_3(t)}{k_2} - \beta \Phi_3(t)] dt \\ \quad + \sigma_2 \Psi_2(t) dB_2(t) + \int_{\mathbb{Y}} \gamma_2(u) \Psi_2(t^-) \tilde{N}(dt, du), \\ d\Psi_3(t) = \Psi_3(t)[\beta \Psi_2(t) + a_3 - \frac{c_3 \Phi_2(t)}{k_2} - \frac{c_3 \Psi_3(t)}{k_2}] dt \\ \quad + \sigma_3 \Psi_3(t) dB_3(t) + \int_{\mathbb{Y}} \gamma_3(u) \Psi_3(t^-) \tilde{N}(dt, du). \end{cases} \tag{6}$$

According to the comparison theorem for stochastic differential equations in [18], we can get that, for  $t \in [0, \tau_e)$ ,

$$\Psi_1(t) \leq x(t) \leq \Phi_1(t), \quad \Psi_2(t) \leq y(t) \leq \Phi_2(t), \quad \Psi_3(t) \leq z(t) \leq \Phi_3(t), \quad \text{a.s.}$$

By virtue of Lemma 4.2 in [10], we can see that  $\Phi_i(t)$  and  $\Psi_i(t)$  ( $i = 1, 2, 3$ ) will not explode in any finite time, which means that  $\Phi_i(t)$  and  $\Psi_i(t)$  are existent on  $t \geq 0$ . Thus, we obtain  $\tau_e = +\infty$ , the proof of Theorem 1 is completed.  $\square$

According to Lemma 4.4 in [10] and equations (5), we know that

$$\limsup_{t \rightarrow \infty} \frac{\ln \Phi_i(t)}{\ln t} \leq 1, \quad \text{a.s.}$$

for  $i = 1, 2, 3$ , when

$$\sup_{t \geq 0} \int_0^t \int_{\mathbb{Y}} \exp(s-t) [\gamma_i(u) - \ln(1 + \gamma_i(u))] \lambda(du) ds < \infty.$$

Noting the limit  $\lim_{t \rightarrow \infty} \frac{\ln t}{t} = 0$ ,  $\limsup_{t \rightarrow \infty} \frac{\ln \Phi_i(t)}{t} \leq 0$  is obtained for  $i = 1, 2, 3$ , a.s. It then follows from the comparison theorem of [18] that

$$\begin{aligned} \limsup_{t \rightarrow \infty} \frac{\ln x(t)}{t} &\leq \limsup_{t \rightarrow \infty} \frac{\ln \Phi_1(t)}{t} \leq 0, \quad \text{a.s.}, \\ \limsup_{t \rightarrow \infty} \frac{\ln y(t)}{t} &\leq \limsup_{t \rightarrow \infty} \frac{\ln \Phi_2(t)}{t} \leq 0, \quad \text{a.s.}, \\ \limsup_{t \rightarrow \infty} \frac{\ln z(t)}{t} &\leq \limsup_{t \rightarrow \infty} \frac{\ln \Phi_3(t)}{t} \leq 0, \quad \text{a.s.} \end{aligned}$$

Thus, we have a lemma as follows.

**Lemma 1** *Assume that, for any  $t \geq 0$  and  $i = 1, 2, 3$ ,*

$$\sup_{t \geq 0} \int_0^t \int_{\mathbb{Y}} \exp(s-t) [\gamma_i(u) - \ln(1 + \gamma_i(u))] \lambda(du) ds < \infty.$$

*Then, for any initial value  $(x_0, y_0, z_0) \in \mathbb{R}_+^3$ , the solution  $(x(t), y(t), z(t))$  of system (3) satisfies*

$$\limsup_{t \rightarrow \infty} \frac{\ln x(t)}{t} \leq 0, \quad \limsup_{t \rightarrow \infty} \frac{\ln y(t)}{t} \leq 0, \quad \limsup_{t \rightarrow \infty} \frac{\ln z(t)}{t} \leq 0, \quad \text{a.s.}$$

### 3 Persistence in the mean and extinction

**Lemma 2** (see Liu et al. [14]) *Suppose that  $Z(t) \in C(\Omega \times [0, \infty), \mathbb{R}_+)$ , and let  $\int_{\mathbb{Y}} [\ln(1 + \gamma_i(u))]^2 \lambda(du) < c$ ,  $i = 1, 2$ , hold.*

(I) *If there exist two positive constants  $T$  and  $\delta_0$  such that*

$$\ln Z(t) \leq \delta t - \delta_0 \int_0^t Z(s) ds + \alpha B(t) + \sum_{i=1}^2 \delta_i \int_0^t \int_{\mathbb{Y}} \ln(1 + \gamma_i(u)) \tilde{N}(ds, du), \quad \text{a.s.}$$

for all  $t \geq T$ , where  $\alpha, \delta_1$  and  $\delta_2$  are constants, then

$$\begin{cases} \limsup_{t \rightarrow \infty} \frac{1}{t} \int_0^t Z(s) ds \leq \delta/\delta_0, & \text{a.s., if } \delta \geq 0; \\ \lim_{t \rightarrow \infty} Z(t) = 0, & \text{a.s., if } \delta < 0. \end{cases}$$

(II) If there exist three positive constants  $T, \delta$  and  $\delta_0$  such that

$$\ln Z(t) \geq \delta t - \delta_0 \int_0^t Z(s) ds + \alpha B(t) + \sum_{i=1}^2 \delta_i \int_0^t \int_{\mathbb{Y}} \ln(1 + \gamma_i(u)) \tilde{N}(ds, du), \quad \text{a.s.}$$

for all  $t \geq T$ , then  $\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t Z(s) ds \geq \delta/\delta_0$ , a.s.

**Definition 1** Assume  $m(t)$  is the density of a population at time  $t$ :

- (1) If  $\lim_{t \rightarrow \infty} m(t) = 0$ , a.s., then species  $m(t)$  is said to be extinct.
- (2) If  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t m(s) ds > 0$ , a.s., then species  $m(t)$  is said to be stable in the mean.
- (3) If  $\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t m(s) ds > 0$ , a.s., then species  $m(t)$  is said to be strongly persistent in the mean.
- (4) If  $\limsup_{t \rightarrow \infty} \frac{1}{t} \int_0^t m(s) ds > 0$ , a.s., then species  $m(t)$  is said to be weakly persistent in the mean.

Obviously, if species  $m(t)$  is stable in the mean, it must be strongly or weakly persistent in the mean.

Making use of Itô's formula to system (3) gives

$$\begin{cases} d \ln x(t) = [w_1 - b_1 x(t) - \frac{c_1 y(t)}{x(t)+k_1} - \frac{\bar{p}c_1 z(t)}{x(t)+k_1}] dt \\ \quad + \sigma_1 dB_1(t) + \int_{\mathbb{Y}} \ln(1 + \gamma_1(u)) \tilde{N}(dt, du), \\ d \ln y(t) = [w_2 - \frac{c_2(y(t)+z(t))}{x(t)+k_2} - \beta z(t)] dt + \sigma_2 dB_2(t) + \int_{\mathbb{Y}} \ln(1 + \gamma_2(u)) \tilde{N}(dt, du), \\ d \ln z(t) = [w_3 + \beta y(t) - \frac{c_3(y(t)+z(t))}{x(t)+k_2}] dt + \sigma_3 dB_3(t) + \int_{\mathbb{Y}} \ln(1 + \gamma_3(u)) \tilde{N}(dt, du). \end{cases} \quad (7)$$

For both sides of system (7), integrating from 0 to  $t$  and multiplying by  $\frac{1}{t}$ , (7) equals

$$\begin{cases} \frac{\ln x(t) - \ln x_0}{t} = w_1 - b_1 \frac{1}{t} \int_0^t x(s) ds - \frac{1}{t} \int_0^t \frac{c_1 y(s)}{x(s)+k_1} ds - \frac{1}{t} \int_0^t \frac{\bar{p}c_1 z(s)}{x(s)+k_1} ds + \frac{\sigma_1 B_1(t)}{t} + \frac{M_1(t)}{t}, \\ \frac{\ln y(t) - \ln y_0}{t} = w_2 - \frac{1}{t} \int_0^t \frac{c_2(y(s)+z(s))}{x(s)+k_2} ds - \beta \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_2 B_2(t)}{t} + \frac{M_2(t)}{t}, \\ \frac{\ln z(t) - \ln z_0}{t} = w_3 + \beta \frac{1}{t} \int_0^t y(s) ds - \frac{1}{t} \int_0^t \frac{c_3(y(s)+z(s))}{x(s)+k_2} ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}, \end{cases} \quad (8)$$

where  $M_i(t) = \int_0^t \int_{\mathbb{Y}} \ln(1 + \gamma_i(u)) \tilde{N}(ds, du)$ ,  $i = 1, 2, 3$ .

$M_i(t)$  are local martingale, by Proposition 2.4 of [19] and Assumption 1, its quadratic variations are

$$\begin{aligned} \langle M_i, M_i \rangle(t) &= \int_0^t \int_{\mathbb{Y}} [\ln(1 + \gamma_i(u))]^2 \lambda(du) ds \\ &= t \int_{\mathbb{Y}} [\ln(1 + \gamma_i(u))]^2 \lambda(du) \\ &< ct, \quad i = 1, 2, 3. \end{aligned}$$

According to the law of strong large numbers, we have

$$\lim_{t \rightarrow \infty} \frac{M_i(t)}{t} = 0, \quad \text{a.s. } i = 1, 2, 3. \quad (9)$$

Obviously,

$$\lim_{t \rightarrow \infty} \frac{B_i(t)}{t} = 0, \quad \text{a.s. } i = 1, 2, 3. \tag{10}$$

In the following, we discuss the persistence in the mean and extinction for prey and predator populations.

**Theorem 2**

- (i) If  $w_i < 0$  ( $i = 1, 2, 3$ ), the prey population  $x(t)$  and both the predator populations  $y(t)$  and  $z(t)$  will be extinct with probability 1.
- (ii) If  $w_1 < 0, w_2 < 0$  and  $w_3 > 0$ , the prey population  $x(t)$  and the susceptible predator population  $y(t)$  will go to extinction, a.s., the infected predator population  $z(t)$  will be stable in the mean, i.e.,  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t z(s) ds = \frac{w_3 k_2}{c_3}$ .
- (iii) If  $w_1 < 0, w_2 > 0$  and  $w_3 < 0$ , the prey population  $x(t)$  will go to extinction, the growth of the predator population is divided into the following two cases:
  - (a) when  $w_3 + \frac{\beta k_2}{c_2} w_2 < 0$ , then the susceptible predator  $y(t)$  will be stable in the mean,  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t y(s) ds = \frac{w_2 k_2}{c_2}$ , and the infected predator  $z(t)$  will be extinct, a.s.
  - (b) when  $w_3 + \frac{\beta k_2}{c_2} w_2 > 0, w_2 > \frac{(c_2 + k_2 \beta)(w_3 c_2 + \beta k_2 w_2)}{c_2 c_3}$  and  $w_3 + \beta Q > \frac{c_3 w_2}{c_2}$ , where  $Q = \frac{k_2}{c_2} [w_2 - \frac{(c_2 + k_2 \beta)(w_3 c_2 + \beta k_2 w_2)}{c_2 c_3}]$ , the whole predator population  $y(t)$  and  $z(t)$  will be strongly persistent in the mean, a.s.
- (iv) If  $w_1 > 0, w_2 < 0$  and  $w_3 < 0$ , then the prey population  $x(t)$  will be stable in the mean, namely  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t x(s) ds = \frac{w_1}{b_1}$ , a.s. Both the susceptible predator  $y(t)$  and the infected predator  $z(t)$  will die out, a.s.
- (v) Assume that, for any  $t \geq 0$  and  $i = 1, 2, 3$ ,

$$\sup_{t \geq 0} \int_0^t \int_{\mathbb{Y}} \exp(s - t) [\gamma_i(u) - \ln(1 + \gamma_i(u))] \lambda(du) ds < \infty,$$

then if  $w_1 > \frac{(w_2 + w_3) k_2 c_1}{(c_2 + c_3 + k_2 \beta) k_1}, w_2 > 0$  and  $w_3 > 0$ , then all the populations will be strongly persistent in the mean, where  $\liminf_{t \rightarrow \infty} [\frac{1}{t} \int_0^t x(s) ds] \geq \frac{w_1}{b_1} - \frac{(w_2 + w_3) k_2 c_1}{b_1 k_1 (c_2 + c_3 + k_2 \beta)}$  and  $\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t [y(s) + z(s)] ds \geq \frac{(w_2 + w_3) k_2}{c_2 + c_3 + k_2 \beta}$ , a.s.

*Proof* Making use of (8), we get

$$\begin{cases} \frac{1}{t} \ln \frac{x(t)}{x_0} \leq w_1 - b_1 \frac{1}{t} \int_0^t x(s) ds + \frac{\sigma_1 B_1(t)}{t} + \frac{M_1(t)}{t}, \\ \frac{1}{t} \ln \frac{y(t)}{y_0} \leq w_2 - \frac{1}{t} \int_0^t \frac{c_2 y(s)}{x(s) + k_2} ds + \frac{\sigma_2 B_2(t)}{t} + \frac{M_2(t)}{t}, \\ \frac{1}{t} \ln \frac{z(t)}{z_0} \leq w_3 + \beta \frac{1}{t} \int_0^t y(s) ds - \frac{1}{t} \int_0^t \frac{c_3 z(s)}{x(s) + k_2} ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}. \end{cases} \tag{11}$$

(i) According to (11),

$$\frac{1}{t} \ln \frac{x(t)}{x_0} \leq w_1 - b_1 \frac{1}{t} \int_0^t x(s) ds + \frac{\sigma_1 B_1(t)}{t} + \frac{M_1(t)}{t},$$

noting (9) and (10), by virtue of Lemma 2, we can get  $\lim_{t \rightarrow \infty} x(t) = 0$  when  $w_1 < 0$ . Then, for an arbitrary small constant  $\varepsilon_1$  which satisfies  $0 < \varepsilon_1 < -\frac{w_3}{\beta}$ , there exists a sufficiently

large number  $T_1$  ( $T_1 > 0$ ) such that

$$-\varepsilon_1 \leq x(t) \leq \varepsilon_1 \tag{12}$$

for  $t \geq T_1$ ; therefore,

$$\frac{1}{t} \ln \frac{y(t)}{y_0} \leq w_2 - \frac{1}{t} \int_0^t \frac{c_2 y(s)}{\varepsilon_1 + k_2} ds + \frac{\sigma_2 B_2(t)}{t} + \frac{M_2(t)}{t}. \tag{13}$$

Because of  $w_2 < 0$ , making use of Lemma 2,  $\lim_{t \rightarrow \infty} y(t) = 0$ , a.s. In this case, there is a sufficiently large constant  $T_2$  ( $T_2 > T_1$ ) satisfying  $-\varepsilon_1 \leq y(t) \leq \varepsilon_1$ , when  $t \geq T_2$ , a.s. Thus,

$$\begin{aligned} \frac{1}{t} \ln \frac{z(t)}{z_0} &\leq w_3 + \beta \frac{1}{t} \int_0^t y(s) ds - \frac{1}{t} \int_0^t \frac{c_3 z(s)}{x(s) + k_2} ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t} \\ &\leq w_3 + \beta \frac{1}{t} \int_0^t \varepsilon_1 ds - \frac{1}{t} \int_0^t \frac{c_3 z(s)}{\varepsilon_1 + k_2} ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t} \\ &\leq (w_3 + \beta \varepsilon_1) - \frac{c_3}{\varepsilon_1 + k_2} \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}. \end{aligned}$$

Notice  $0 < \varepsilon_1 < -\frac{w_3}{\beta}$ , it follows from Lemma 2 that  $\lim_{t \rightarrow \infty} z(t) = 0$ , a.s.

(ii) Similar to case (i), when  $w_1 < 0$ ,  $w_2 < 0$ , it easily shows that  $\lim_{t \rightarrow \infty} x(t) = 0$ ,  $\lim_{t \rightarrow \infty} y(t) = 0$ , and for an arbitrary small constant  $\bar{\varepsilon}_1 > 0$ ,

$$\frac{1}{t} \ln \frac{z(t)}{z_0} \leq (w_3 + \beta \bar{\varepsilon}_1) - \frac{c_3}{\bar{\varepsilon}_1 + k_2} \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}.$$

Making use of Lemma 2,  $w_3 > 0$  and the arbitrariness of  $\bar{\varepsilon}_1$ ,  $\limsup_{t \rightarrow \infty} \frac{1}{t} \int_0^t z(s) ds \leq \frac{w_3 k_2}{c_3}$ , a.s. is obviously verified. In contrast,

$$\frac{1}{t} \ln \frac{z(t)}{z_0} \geq w_3 - \frac{c_3}{k_2} \frac{1}{t} \int_0^t y(s) ds - \frac{c_3}{k_2} \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}.$$

Noting that  $\lim_{t \rightarrow \infty} y(t) = 0$ , a.s. and Lemma 2, we have

$$\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t z(s) ds \geq \frac{w_3 k_2}{c_3}, \quad \text{a.s.,}$$

then we obtain  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t z(s) ds = \frac{w_3 k_2}{c_3}$ , a.s.

(iii) It is easy to show that  $\lim_{t \rightarrow \infty} x(t) = 0$ , a.s. when  $w_1 < 0$ . On the basis of (13) and Lemma 2, for  $t > T_1$ ,  $\limsup_{t \rightarrow \infty} \frac{1}{t} \int_0^t y(s) ds \leq \frac{w_2(\varepsilon_1 + k_2)}{c_2}$ , a.s. under the condition of  $w_2 > 0$ . Considering the arbitrariness of  $\varepsilon_1$ ,  $\limsup_{t \rightarrow \infty} \frac{1}{t} \int_0^t y(s) ds \leq \frac{w_2 k_2}{c_2}$ , a.s. Therefore, for an arbitrary constant  $\varepsilon_3 > 0$  ( $0 < \varepsilon_3 < \frac{1}{\beta} |w_3 + \frac{\beta k_2}{c_2} w_2|$ ), there exists a number  $T_3$  ( $T_3 > T_1$ ) such that

$$\frac{1}{t} \int_0^t y(s) ds \leq \frac{w_2 k_2}{c_2} + \varepsilon_3, \quad \text{a.s.}$$

for  $t > T_3$ . Making use of (12), then for  $t > T_3$ , we have

$$\begin{aligned} \frac{1}{t} \ln \frac{z(t)}{z_0} &\leq w_3 + \beta \frac{1}{t} \int_0^t y(s) ds - \frac{1}{t} \int_0^t \frac{c_3 z(s)}{\varepsilon_1 + k_2} ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t} \\ &\leq w_3 + \beta \left( \frac{w_2 k_2}{c_2} + \varepsilon_3 \right) - \frac{c_3}{\varepsilon_1 + k_2} \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}. \end{aligned}$$

(a) If  $w_3 + \frac{\beta k_2}{c_2} w_2 < 0$ , then

$$\lim_{t \rightarrow \infty} z(t) = 0, \quad \text{a.s.} \tag{14}$$

For the susceptible predator population,

$$\frac{1}{t} \ln \frac{y(t)}{y_0} \geq w_2 - \frac{1}{t} \int_0^t \frac{c_2 y(s)}{k_2} ds - \frac{1}{t} \int_0^t \frac{c_2 z(s)}{k_2} ds - \beta \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_2 B_2(t)}{t} + \frac{M_2(t)}{t}.$$

Because of (9), (10), (14) and Lemma 2, we have  $\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t y(s) ds \geq \frac{w_2 k_2}{c_2}$ , a.s. Then  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t y(s) ds = \frac{w_2 k_2}{c_2}$ , a.s.

(b) If  $w_3 + \frac{\beta k_2}{c_2} w_2 > 0$ , by virtue of Lemma 2 and the arbitrariness of  $\varepsilon_1$  and  $\varepsilon_3$ , we have

$$\limsup_{t \rightarrow \infty} \frac{1}{t} \int_0^t z(s) ds \leq \frac{(w_3 c_2 + \beta k_2 w_2) k_2}{c_2 c_3}, \quad \text{a.s.} \tag{15}$$

Then

$$\frac{1}{t} \ln \frac{y(t)}{y_0} \geq w_2 - \frac{c_2}{k_2} \frac{1}{t} \int_0^t y(s) ds - \left( \frac{c_2}{k_2} + \beta \right) \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_2 B_2(t)}{t} + \frac{M_2(t)}{t}.$$

Notice (15), for sufficiently large  $t$ ,  $\frac{1}{t} \int_0^t z(s) ds \leq \frac{(w_3 c_2 + \beta k_2 w_2) k_2}{c_2 c_3} + \varepsilon_4$ , a.s., where  $\varepsilon_4$  is an arbitrary positive number. Hence, by (9), (10), Lemma 2 and the arbitrariness of  $\varepsilon_4$ , we have

$$\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t y(s) ds \geq \frac{k_2}{c_2} \left[ w_2 - \frac{(c_2 + k_2 \beta)(w_3 c_2 + \beta k_2 w_2)}{c_2 c_3} \right],$$

when  $w_2 > \frac{(c_2 + k_2 \beta)(w_3 c_2 + \beta k_2 w_2)}{c_2 c_3}$ . Here we define  $Q = \frac{k_2}{c_2} \left[ w_2 - \frac{(c_2 + k_2 \beta)(w_3 c_2 + \beta k_2 w_2)}{c_2 c_3} \right]$ .

In this case,

$$\frac{1}{t} \ln \frac{z(t)}{z_0} \geq w_3 + \beta \frac{1}{t} \int_0^t y(s) ds - \frac{c_3}{k_2} \frac{1}{t} \int_0^t y(s) ds - \frac{c_3}{k_2} \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}.$$

For sufficiently large  $t$ , we have

$$\frac{1}{t} \ln \frac{z(t)}{z_0} \geq \left[ w_3 + \beta Q - \frac{c_3}{k_2} \frac{w_2 k_2}{c_2} \right] - \frac{c_3}{k_2} \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}.$$

If  $w_3 + \beta Q > \frac{c_3 w_2}{c_2}$ , then by Lemma 2, we get

$$\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t z(s) ds \geq \frac{[w_3 + \beta Q - \frac{c_3 w_2}{c_2}] k_2}{c_3}, \quad \text{a.s.}$$

(iv) According to (11), we have

$$\frac{1}{t} \ln \frac{x(t)}{x_0} \leq w_1 - b_1 \frac{1}{t} \int_0^t x(s) ds + \frac{\sigma_1 B_1(t)}{t} + \frac{M_1(t)}{t}.$$

Making use of Lemma 2, (9) and (10), we have

$$\limsup_{t \rightarrow \infty} \frac{1}{t} \int_0^t x(s) ds \leq \frac{w_1}{b_1}, \quad \text{a.s.} \tag{16}$$

and

$$\frac{1}{t} \ln \frac{y(t)}{y_0} \leq w_2 + \frac{\sigma_2 B_2(t)}{t} + \frac{M_2(t)}{t}.$$

Because of  $w_2 < 0$  and (9), (10), we get  $\limsup_{t \rightarrow \infty} \frac{\ln y(t)}{t} < 0$ , a.s., namely

$$\lim_{t \rightarrow \infty} y(t) = 0, \quad \text{a.s.} \tag{17}$$

Then

$$\frac{1}{t} \ln \frac{z(t)}{z_0} \leq w_3 + \beta \frac{1}{t} \int_0^t y(s) ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}.$$

Considering (9), (10), (17) and  $w_3 < 0$ , we have  $\limsup_{t \rightarrow \infty} \frac{\ln z(t)}{t} < 0$ ; therefore,

$$\lim_{t \rightarrow \infty} z(t) = 0, \quad \text{a.s.} \tag{18}$$

At this time, for sufficiently large  $t$ ,

$$\frac{1}{t} \ln \frac{x(t)}{x_0} \geq w_1 - b_1 \frac{1}{t} \int_0^t x(s) ds - \frac{c_1}{k_1} \frac{1}{t} \int_0^t y(s) ds - \frac{\tilde{p}c_1}{k_1} \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_1 B_1(t)}{t} + \frac{M_1(t)}{t}.$$

By (17), (18) and Lemma 2,

$$\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t x(s) ds \geq \frac{w_1}{b_1}, \quad \text{a.s.} \tag{19}$$

Combining (19) with (16), we have  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t x(s) ds = \frac{w_1}{b_1}$ , a.s.

(v) It is easy to have  $\limsup_{t \rightarrow \infty} \frac{1}{t} \int_0^t x(s) ds \leq \frac{w_1}{b_1}$ , a.s. when  $w_1 > 0$  and

$$\begin{aligned} \frac{1}{t} \ln \frac{y(t)}{y_0} &\geq w_2 - \frac{c_2}{k_2} \frac{1}{t} \int_0^t y(s) ds - \left( \frac{c_2}{k_2} + \beta \right) \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_2 B_2(t)}{t} + \frac{M_2(t)}{t}, \\ \frac{1}{t} \ln \frac{z(t)}{z_0} &\geq w_3 - \frac{c_3}{k_2} \frac{1}{t} \int_0^t y(s) ds - \frac{c_3}{k_2} \frac{1}{t} \int_0^t z(s) ds + \frac{\sigma_3 B_3(t)}{t} + \frac{M_3(t)}{t}. \end{aligned}$$

Then

$$\begin{aligned} \frac{1}{t} \ln \frac{y(t)}{y_0} + \frac{1}{t} \ln \frac{z(t)}{z_0} &\geq (w_2 + w_3) - \frac{c_2 + c_3}{k_2} \frac{1}{t} \int_0^t y(s) ds - \frac{c_2 + c_3 + k_2 \beta}{k_2} \frac{1}{t} \int_0^t z(s) ds \\ &\quad + \sum_{i=2}^3 \left[ \frac{\sigma_i B_i(t)}{t} + \frac{M_i(t)}{t} \right], \end{aligned}$$

namely

$$\begin{aligned} \frac{1}{t} \ln(y(t)z(t)) - \frac{1}{t} \ln(y_0z_0) &\geq (w_2 + w_3) - \frac{c_2 + c_3 + k_2\beta}{k_2} \frac{1}{t} \int_0^t [y(s) + z(s)] ds \\ &\quad + \sum_{i=2}^3 \left[ \frac{\sigma_i B_i(t)}{t} + \frac{M_i(t)}{t} \right]. \end{aligned}$$

By virtue of the formula  $(a + b)^2 \geq 2ab$  for  $a, b \in R_+$ , we can get  $2 \ln(a + b) \geq \ln 2 + \ln(ab)$ , hence

$$\begin{aligned} \frac{2 \ln(y(t) + z(t))}{t} - \frac{\ln 2}{t} - \frac{\ln(y_0z_0)}{t} &\geq (w_2 + w_3) - \frac{(c_2 + c_3 + k_2\beta)}{k_2} \frac{1}{t} \int_0^t [y(s) + z(s)] ds \\ &\quad + \sum_{i=2}^3 \left[ \frac{\sigma_i B_i(t)}{t} + \frac{M_i(t)}{t} \right], \end{aligned}$$

then

$$\begin{aligned} \frac{\ln(y(t) + z(t))}{t} - \frac{\ln 2}{2t} - \frac{\ln(y_0z_0)}{2t} &\geq \frac{(w_2 + w_3)}{2} - \frac{(c_2 + c_3 + k_2\beta)}{2k_2} \frac{1}{t} \int_0^t [y(s) + z(s)] ds \\ &\quad + \frac{1}{2} \sum_{i=2}^3 \left[ \frac{\sigma_i B_i(t)}{t} + \frac{M_i(t)}{t} \right]. \end{aligned}$$

By Lemma 2, we get

$$\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t [y(s) + z(s)] ds \geq \frac{(w_2 + w_3)k_2}{c_2 + c_3 + k_2\beta}, \quad \text{a.s.}$$

In this case, considering  $0 < \tilde{p} < 1$ , we have

$$\frac{1}{t} \ln \frac{x(t)}{x_0} \geq w_1 - b_1 \frac{1}{t} \int_0^t x(s) ds - \frac{c_1}{k_1} \frac{1}{t} \int_0^t [y(s) + z(s)] ds + \frac{\sigma_1 B_1(t)}{t} + \frac{M_1(t)}{t}.$$

Taking the inferior limit, by virtue of Lemma 1 and (9), (10), we get

$$\begin{aligned} 0 &\geq \limsup_{t \rightarrow \infty} \frac{1}{t} \ln \frac{x(t)}{x_0} \geq \liminf_{t \rightarrow \infty} \frac{1}{t} \ln \frac{x(t)}{x_0} \\ &\geq w_1 - b_1 \liminf_{t \rightarrow \infty} \left[ \frac{1}{t} \int_0^t x(s) ds \right] - \frac{c_1}{k_1} \liminf_{t \rightarrow \infty} \left[ \frac{1}{t} \int_0^t [y(s) + z(s)] ds \right] \\ &\quad + \liminf_{t \rightarrow \infty} \left[ \frac{\sigma_1 B_1(t)}{t} + \frac{M_1(t)}{t} \right]. \end{aligned}$$

Hence,

$$\liminf_{t \rightarrow \infty} \left[ \frac{1}{t} \int_0^t x(s) ds \right] \geq \frac{w_1}{b_1} - \frac{(w_2 + w_3)k_2c_1}{b_1k_1(c_2 + c_3 + k_2\beta)}, \quad \text{a.s.} \quad \square$$

#### 4 Stochastically ultimate boundedness

From the biological point of view, the nonexplosion and positiveness property of solution in a population dynamical system is not enough. In this section we will show that the  $p$ th moment of solution (3) is bounded.

**Definition 2** The solution  $X(t)$  of system (3) is called stochastically ultimate bounded if, for any  $\epsilon \in (0, 1)$ , there is a constant  $H := H(\epsilon)$  such that, for any  $X(0) \in R_+^3$ ,

$$\limsup_{t \rightarrow \infty} P\{|X(t)| \leq H\} \geq 1 - \epsilon.$$

**Theorem 3** Assume  $(x(t), y(t), z(t))^T$  to be a solution to system (3) with the initial value  $(x_0, y_0, z_0) \in R_+^3$ . If  $k_2 > \frac{(1+a_1)^2}{4b_1}$  for all  $p > 0$ , there exist constants  $\bar{K}_i(p)$  ( $i = 1, 2, 3$ ) such that

$$\limsup_{t \rightarrow \infty} E[x^p(t)] \leq \bar{K}_1(p), \quad \limsup_{t \rightarrow \infty} E[y^p(t)] \leq \bar{K}_2(p), \quad \limsup_{t \rightarrow \infty} E[z^p(t)] \leq \bar{K}_3(p).$$

*Proof* Making use of Itô’s formula and then taking expectations yield

$$\begin{aligned} d(e^t x^p(t)) &= e^t x^p(t) \left\{ 1 + p \left[ a_1 - b_1 x - \frac{c_1 y}{x + k_1} - \frac{\tilde{p}c_1 z}{x + k_1} \right] + \frac{p(p-1)}{2} \sigma_1^2 \right. \\ &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_1(u))^p - 1 - p\gamma_1(u)] \lambda(du) \right\} dt + p\sigma_1 e^t x^p(t) dB_1(t) \\ &\quad + \int_{\mathbb{Y}} e^t x^p(t) [(1 + \gamma_1(u))^p - 1] \tilde{N}(dt, du), \\ E(e^t x^p(t)) &= x^p(0) + pE \int_0^t e^s x^p(s) \left\{ \frac{1}{p} + a_1 + \frac{(p-1)}{2} \sigma_1^2 \right. \\ &\quad \left. + \frac{1}{p} \int_{\mathbb{Y}} [(1 + \gamma_1(u))^p - 1 - p\gamma_1(u)] \lambda(du) \right. \\ &\quad \left. - b_1 x(s) - \frac{c_1 y(s)}{x(s) + k_1} - \frac{\tilde{p}c_1 z(s)}{x(s) + k_1} \right\} ds \\ &\leq x^p(0) + pE \int_0^t e^s x^p(s) \left\{ \frac{1}{p} + a_1 + \frac{(p-1)}{2} \sigma_1^2 \right. \\ &\quad \left. + \frac{1}{p} \int_{\mathbb{Y}} [(1 + \gamma_1(u))^p - 1 - p\gamma_1(u)] \lambda(du) - b_1 x(s) \right\} ds. \end{aligned}$$

Note for a function  $f(v) = v^p(a - bv)$ ,  $v \in [0, +\infty)$ ,  $p > 0$ ,  $a > 0$ ,  $b > 0$ ,  $f(v)$  has a maximum value at  $\frac{ap}{(p+1)b}$ , namely

$$f(v) \leq f\left(\frac{ap}{(p+1)b}\right) = \frac{a^{p+1} p^p}{(p+1)^{p+1} b^p}. \tag{20}$$

Therefore,

$$E(e^t x^p(t)) \leq x^p(0) + E \int_0^t e^s \bar{K}_1(p) ds \leq x^p(0) + \bar{K}_1(p)(e^t - 1),$$

where

$$\bar{K}_1(p) = \frac{[1 + a_1 p + \frac{p(p-1)}{2} \sigma_1^2 + \int_{\mathbb{Y}} [(1 + \gamma_1(u))^p - 1 - p\gamma_1(u)] \lambda(du)]^{p+1}}{(p+1)^{p+1} b_1^p}.$$

Let  $t \rightarrow +\infty$ , we have

$$\limsup_{t \rightarrow \infty} E[x^p(t)] \leq \bar{K}_1(p), \quad \text{a.s.} \tag{21}$$

Similarly, we can compute

$$\begin{aligned} d(e^t y^p(t)) &= e^t y^p(t) \left\{ 1 + p \left[ a_2 - \frac{c_2(y+z)}{x+k_2} - \beta z \right] + \frac{p(p-1)}{2} \sigma_2^2 \right. \\ &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_2(u))^p - p\gamma_2(u) - 1] \lambda(du) \right\} dt + p\sigma_2 e^t y^p(t) dB_2(t) \\ &\quad + \int_{\mathbb{Y}} e^t y^p(t) [(1 + \gamma_2(u))^p - 1] \tilde{N}(dt, du). \end{aligned}$$

Thus, taking expectations leads to

$$\begin{aligned} E(e^t y^p(t)) &= y^p(0) + E \int_0^t e^s y^p(s) \left[ 1 + pa_2 + \frac{p(p-1)}{2} \sigma_2^2 \right. \\ &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_2(u))^p - p\gamma_2(u) - 1] \lambda(du) \right. \\ &\quad \left. - \frac{c_2 p(y(s) + z(s))}{x(s) + k_2} - p\beta z(s) \right] ds \\ &\leq y^p(0) + E \int_0^t e^s y^p(s) \left[ 1 + pa_2 + \frac{p(p-1)}{2} \sigma_2^2 \right. \\ &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_2(u))^p - p\gamma_2(u) - 1] \lambda(du) \right. \\ &\quad \left. - \frac{c_2 p y(s)}{k_2} + \frac{c_2 p x(s) y(s)}{k_2(x(s) + k_2)} \right] ds. \end{aligned}$$

By virtue of (21), for arbitrary small  $\varepsilon_0 > 0$  ( $\varepsilon_0 < k_2 - \bar{K}_1(1)$ ), there exists a positive constant  $T_0$  satisfying

$$E(x^p(t)) \leq \bar{K}_1(p) + \varepsilon_0$$

for  $t > T_0$ . Therefore,

$$\begin{aligned} E(e^t y^p(t)) &\leq y^p(0) + \int_0^t e^s E \left\{ y^p(s) \left[ 1 + pa_2 + \frac{p(p-1)}{2} \sigma_2^2 + \int_{\mathbb{Y}} [(1 + \gamma_2(u))^p \right. \right. \\ &\quad \left. \left. - p\gamma_2(u) - 1] \lambda(du) - \frac{c_2 p y(s)}{k_2} \right] \right\} ds + \frac{c_2 p}{k_2^2} \int_0^t e^s E(x(s)) E(y^{p+1}(s)) ds \\ &\leq y^p(0) + \int_0^t e^s E \left\{ y^p(s) \left[ 1 + pa_2 + \frac{p(p-1)}{2} \sigma_2^2 + \int_{\mathbb{Y}} [(1 + \gamma_2(u))^p \right. \right. \\ &\quad \left. \left. - p\gamma_2(u) - 1] \lambda(du) - \frac{c_2 p y(s)}{k_2} \right] \right\} ds + \frac{c_2 p}{k_2^2} (\bar{K}_1(1) + \varepsilon_0) \int_0^t e^s E(y^{p+1}(s)) ds \\ &\leq y^p(0) + \int_0^t e^s E \left\{ y^p(s) \left[ 1 + pa_2 + \frac{p(p-1)}{2} \sigma_2^2 + \int_{\mathbb{Y}} [(1 + \gamma_2(u))^p \right. \right. \\ &\quad \left. \left. - p\gamma_2(u) - 1] \lambda(du) - \frac{c_2 p y(s)}{k_2} + \frac{c_2 p}{k_2^2} (\bar{K}_1(1) + \varepsilon_0) y(s) \right] \right\} ds. \end{aligned}$$

If  $k_2 > \bar{K}_1(1)$ , then  $k_2 > \bar{K}_1(1) + \varepsilon_0$ . According to (20), the following result is available:

$$E(e^t y^p(t)) \leq y^p(0) + \int_0^t e^s K_2(p) ds = y^p(0) + K_2(p)(e^t - 1).$$

Here

$$K_2(p) = \frac{[1 + pa_2 + \frac{p(p-1)}{2}\sigma_2^2 + \int_{\mathbb{Y}} [(1 + \gamma_2(u))^p - p\gamma_2(u) - 1]\lambda(du)]^{p+1}}{(p + 1)^{p+1}} \cdot \frac{k_2^{2p}}{c_2^p(k_2 - \bar{K}_1(1) - \varepsilon_0)^p},$$

where  $\bar{K}_1(1) = \frac{(1+a_1)^2}{4b_1}$ . Similarly,  $\limsup_{t \rightarrow \infty} E[y^p(t)] \leq K_2(p)$ , here because of the arbitrariness of  $\varepsilon_0$ , define

$$\bar{K}_2(p) = \frac{[1 + pa_2 + \frac{p(p-1)}{2}\sigma_2^2 + \int_{\mathbb{Y}} [(1 + \gamma_2(u))^p - p\gamma_2(u) - 1]\lambda(du)]^{p+1}}{(p + 1)^{p+1}} \cdot \frac{k_2^{2p}}{c_2^p(k_2 - \bar{K}_1(1))^p},$$

thus,

$$\limsup_{t \rightarrow \infty} E[y^p(t)] \leq \bar{K}_2(p). \tag{22}$$

And

$$\begin{aligned} d(e^t z^p(t)) &= e^t z^p(t) \left\{ 1 + p \left[ \beta y + a_3 - \frac{c_3(y+z)}{x+k_2} \right] + \frac{p(p-1)}{2} \sigma_3^2 \right. \\ &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1]\lambda(du) \right\} dt + p\sigma_3 e^t z^p(t) dB_3(t) \\ &\quad + \int_{\mathbb{Y}} e^t z^p(t) [(1 + \gamma_3(u))^p - 1] \tilde{N}(dt, du), \\ E(e^t z^p(t)) &= z^p(0) + E \int_0^t e^s z^p(s) \left[ 1 + pa_3 + \frac{p(p-1)}{2} \sigma_3^2 \right. \\ &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1]\lambda(du) \right. \\ &\quad \left. + p\beta y(s) - \frac{c_3 p(y(s) + z(s))}{x(s) + k_2} \right] ds \\ &\leq z^p(0) + E \int_0^t e^s z^p(s) \left[ 1 + pa_3 + \frac{p(p-1)}{2} \sigma_3^2 \right. \\ &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1]\lambda(du) \right. \\ &\quad \left. + p\beta y(s) - \frac{c_3 pz(s)}{x(s) + k_2} \right] ds \\ &\leq z^p(0) + E \int_0^t p\beta e^s z^p(s) y(s) ds + E \int_0^t e^s z^p(s) \left[ 1 + pa_3 + \frac{p(p-1)}{2} \sigma_3^2 \right. \\ &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1]\lambda(du) - \frac{c_3 pz(s)}{k_2} + \frac{c_3 px(s)z(s)}{k_2(x(s) + k_2)} \right] ds. \end{aligned}$$

On the basis of Young’s inequality, we have

$$\begin{aligned}
 y(t)z^p(t) &\leq \frac{y^{p+1}(t)}{p+1} + \frac{p}{p+1} (z^p(t))^{\frac{p+1}{p}} \\
 &\leq \frac{y^{p+1}(t)}{p+1} + \frac{p}{p+1} z^{p+1}(t).
 \end{aligned}$$

Notice  $\frac{p}{p+1} < 1$ , we have  $y(t)z^p(t) \leq \frac{y^{p+1}(t)}{p+1} + z^{p+1}(t)$ , then

$$\begin{aligned}
 E(e^t z^p(t)) &\leq z^p(0) + p\beta E \int_0^t e^s z^{p+1}(s) ds + p\beta E \int_0^t e^s \frac{1}{p+1} y^{p+1}(s) ds \\
 &\quad + E \int_0^t e^s z^p(s) \left[ 1 + pa_3 + \frac{p(p-1)}{2} \sigma_3^2 \right. \\
 &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1] \lambda(du) - \frac{c_3 p z(s)}{k_2} + \frac{c_3 p x(s) z(s)}{k_2^2} \right] ds \\
 &\leq z^p(0) + \frac{p\beta}{p+1} E \int_0^t e^s y^{p+1}(s) ds + E \int_0^t e^s z^p(s) \left[ 1 + pa_3 + \frac{p(p-1)}{2} \sigma_3^2 \right. \\
 &\quad \left. + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1] \lambda(du) - \left( \frac{c_3 p}{k_2} - p\beta \right) z(s) \right] ds \\
 &\quad + \frac{c_3 p}{k_2^2} \int_0^t e^s E(x(s)) E(z^{p+1}(s)) ds.
 \end{aligned}$$

Considering (22), we can see that there exist positive constants  $\bar{K}_2(p)$  and  $\bar{T}_2$  ( $\bar{T}_2 > T_0$ ), for  $t > \bar{T}_2$ , one can obtain that  $E[y^p(t)] \leq \bar{K}_2(p)$ . Then, for  $t > \bar{T}_2$ ,

$$\begin{aligned}
 E(e^t z^p(t)) &\leq z^p(0) + \frac{p\beta}{p+1} \bar{K}_2(p+1) [e^t - 1] + \int_0^t e^s E \left\{ z^p(s) \left[ 1 + pa_3 + \frac{p(p-1)}{2} \sigma_3^2 \right. \right. \\
 &\quad \left. \left. + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1] \lambda(du) \right. \right. \\
 &\quad \left. \left. - \left( \frac{c_3 p}{k_2} - p\beta - \frac{c_3 p}{k_2^2} (\bar{K}_1(1) + \varepsilon_0) \right) z(s) \right] \right\} ds \\
 &\leq z^p(0) + \frac{p\beta}{p+1} \bar{K}_2(p+1) [e^t - 1] + \int_0^t e^s K_3(p) ds \\
 &\leq z^p(0) + \frac{p\beta}{p+1} \bar{K}_2(p+1) [e^t - 1] + K_3(p) [e^t - 1],
 \end{aligned}$$

where

$$\begin{aligned}
 K_3(p) &= \frac{[1 + pa_3 + \frac{p(p-1)}{2} \sigma_3^2 + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1] \lambda(du)]^{p+1}}{(p+1)^{p+1}} \\
 &\quad \cdot \frac{p^p}{[\frac{c_3 p}{k_2} - p\beta - \frac{c_3 p}{k_2^2} (\bar{K}_1(1) + \varepsilon_0)]^p}.
 \end{aligned}$$

Take the superior limit results in

$$\limsup_{t \rightarrow \infty} E[z^p(t)] \leq \frac{p\beta}{p+1} \bar{K}_2(p+1) + K_3(p).$$

Notice the arbitrariness of  $\varepsilon_0$ , let

$$\bar{K}_3(p) = \frac{p\beta}{p+1} \bar{K}_2(p+1) + \frac{[1 + pa_3 + \frac{p(p-1)}{2}\sigma_3^2 + \int_{\mathbb{Y}} [(1 + \gamma_3(u))^p - p\gamma_3(u) - 1]\lambda(du)]^{p+1}}{(p+1)^{p+1}} \cdot \frac{p^p}{[\frac{c_3p}{k_2} - p\beta - \frac{c_3p}{k_2} \bar{K}_1(1)]^p},$$

clearly,

$$\limsup_{t \rightarrow \infty} E[z^p(t)] \leq \bar{K}_3(p),$$

thus the proof of Theorem 3 is completed. □

By virtue of the formula of Lemma 2.1 in [10], we have

$$n^{(1-\frac{p}{2}) \wedge 0} |x|^p \leq \sum_{i=1}^n x_i^p \leq n^{(1-\frac{p}{2}) \vee 0} |x|^p$$

for  $\forall p > 0, x \in R_+^n$ , where  $R_+^n := \{x \in R^n, x_i > 0, 1 \leq i \leq n\}$ .

For solution  $X(t) = (x(t), y(t), z(t))^T \in R_+^3$  and  $p > 0$ , we have

$$3^{(1-\frac{p}{2}) \wedge 0} |X(t)|^p \leq x^p(t) + y^p(t) + z^p(t).$$

By Theorem 3,

$$\limsup_{t \rightarrow \infty} E[X(t)]^p \leq \left(\frac{1}{3}\right)^{(1-\frac{p}{2}) \wedge 0} \limsup_{t \rightarrow \infty} E[x^p(t) + y^p(t) + z^p(t)] \leq \left(\frac{1}{3}\right)^{(1-\frac{p}{2}) \wedge 0} \sum_{i=1}^3 \bar{K}_i(p).$$

As an application of Theorem 3, together with the Chebyshev inequality, we can also establish the following corollary.

**Corollary 1** *Under conditions of Theorem 3, the solution of system (3) is stochastically ultimate bounded.*

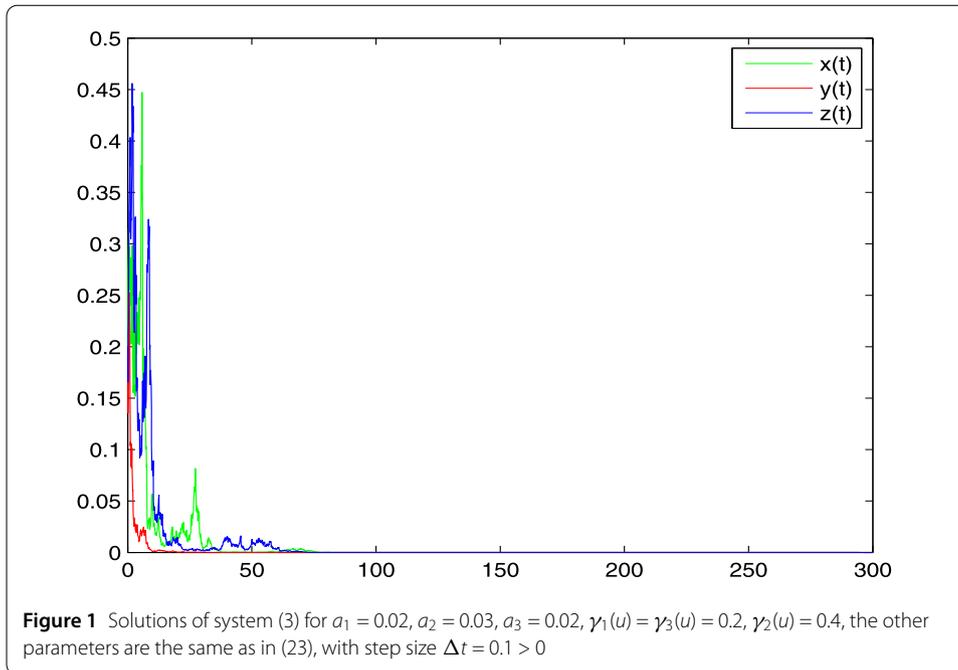
### 5 Numerical simulation

In this section, we give some examples and numerical simulations to illustrate our analytical findings. These numerical simulations are given by the Euler scheme of [20].

For system (3), we choose the initial value  $(x_0, y_0, z_0) = (0.2, 0.1, 0.2)$ , and

$$\begin{aligned} b_1 = 0.3, & \quad c_1 = 0.4, & \quad \tilde{p} = 0.1, & \quad c_2 = 0.6, & \quad k_1 = 0.5, \\ k_2 = 0.5, & \quad \beta = 0.4, & \quad c_3 = 0.5, & \quad \sigma_i = 0.5 \quad (i = 1, 2, 3), \end{aligned} \tag{23}$$

$\mathbb{Y} = (0, +\infty), \lambda(\mathbb{Y}) = 1$ . In the following, we change the values of part parameters to observe the asymptotic behavior of solution for system (3).



- (1) In Figure 1, let  $a_1 = 0.02, a_2 = 0.03, a_3 = 0.02, \gamma_1(u) = \gamma_3(u) = 0.2, \gamma_2(u) = 0.4$ , the other parameters are the same as above, then

$$w_1 = a_1 - \frac{\sigma_1^2}{2} - \int_{\mathbb{Y}} [\gamma_1(u) - \ln(1 + \gamma_1(u))] \lambda(du) = -0.1227 < 0,$$

$$w_2 = a_2 - \frac{\sigma_2^2}{2} - \int_{\mathbb{Y}} [\gamma_2(u) - \ln(1 + \gamma_2(u))] \lambda(du) = -0.1585 < 0,$$

$$w_3 = a_3 - \frac{\sigma_3^2}{2} - \int_{\mathbb{Y}} [\gamma_3(u) - \ln(1 + \gamma_3(u))] \lambda(du) = -0.1227 < 0.$$

Thus, in accordance with case (i) in Theorem 2, all the species will go to extinction, Figure 1 confirms it.

- (2) In Figure 2, let  $a_1 = 0.02, a_2 = 0.03, a_3 = 0.5, \gamma_1(u) = \gamma_2(u) = 0.2, \gamma_3(u) = 0.4$ , the other parameters remain the same as in (23), then

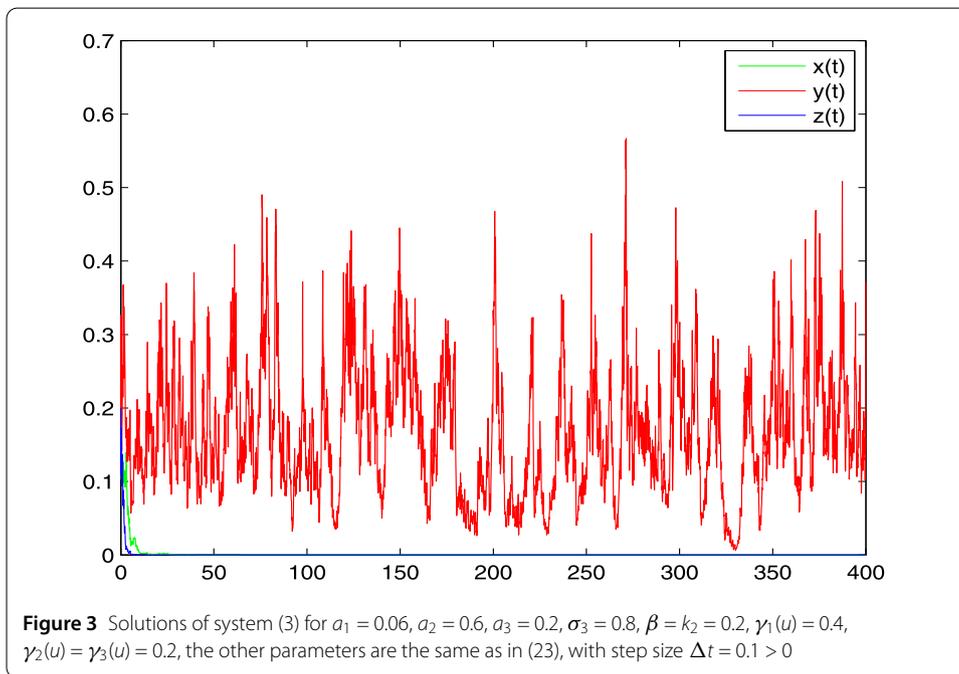
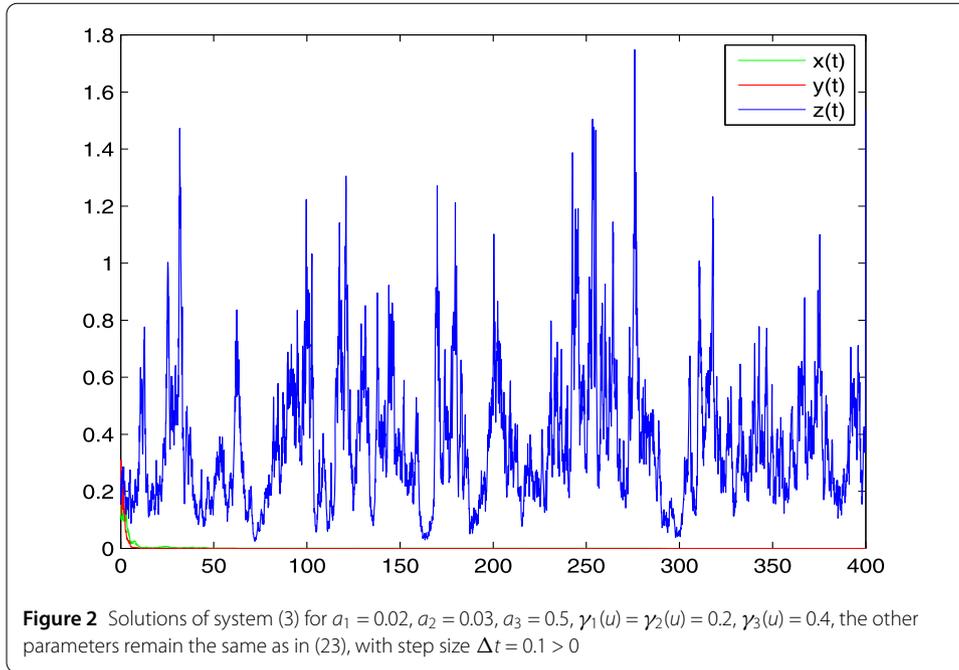
$$w_1 = -0.1227 < 0, \quad w_2 = -0.1127 < 0, \quad w_3 = 0.3115 > 0.$$

By the case (ii) of Theorem 2, the prey population and the susceptible predator population will be extinct, the infected predator population will be stable in the mean. Figure 2 shows it.

- (3) In Figure 3, we assume  $a_1 = 0.06, a_2 = 0.6, a_3 = 0.2, \sigma_3 = 0.8, \beta = k_2 = 0.2, \gamma_1(u) = 0.4, \gamma_2(u) = \gamma_3(u) = 0.2$ , the other parameters are the same as in (23), at this time

$$w_1 = -0.1285 < 0, \quad w_2 = 0.4573 > 0,$$

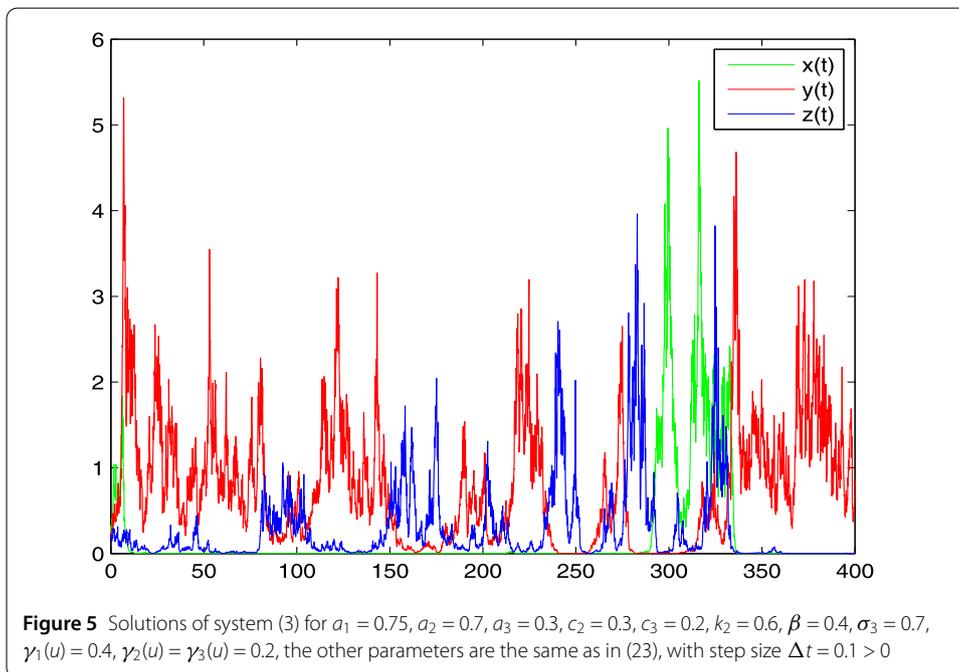
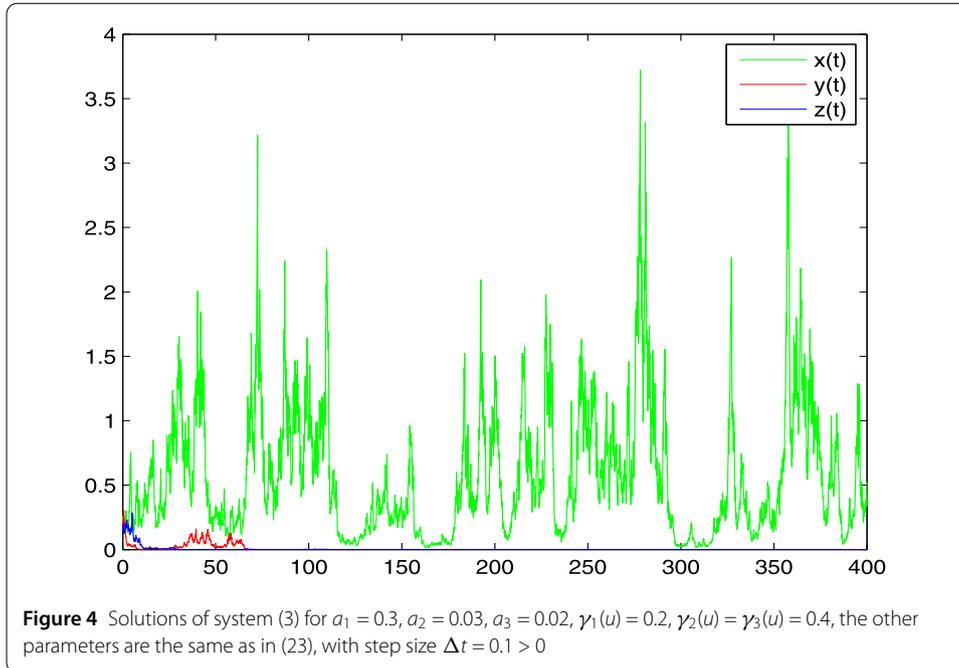
$$w_3 = -0.1377 < 0, \quad w_3 + \frac{\beta k_2}{c_2} w_2 = -0.1072 < 0.$$



By the condition of (iii)(a), the prey population  $x(t)$  and the infected population  $z(t)$  will go to extinction, the susceptible predator population  $y(t)$  will be stable in the mean.

- (4) In Figure 4, we choose  $a_1 = 0.3, a_2 = 0.03, a_3 = 0.02, \gamma_1(u) = 0.2, \gamma_2(u) = \gamma_3(u) = 0.4$ , the other parameters are the same as in (23), then

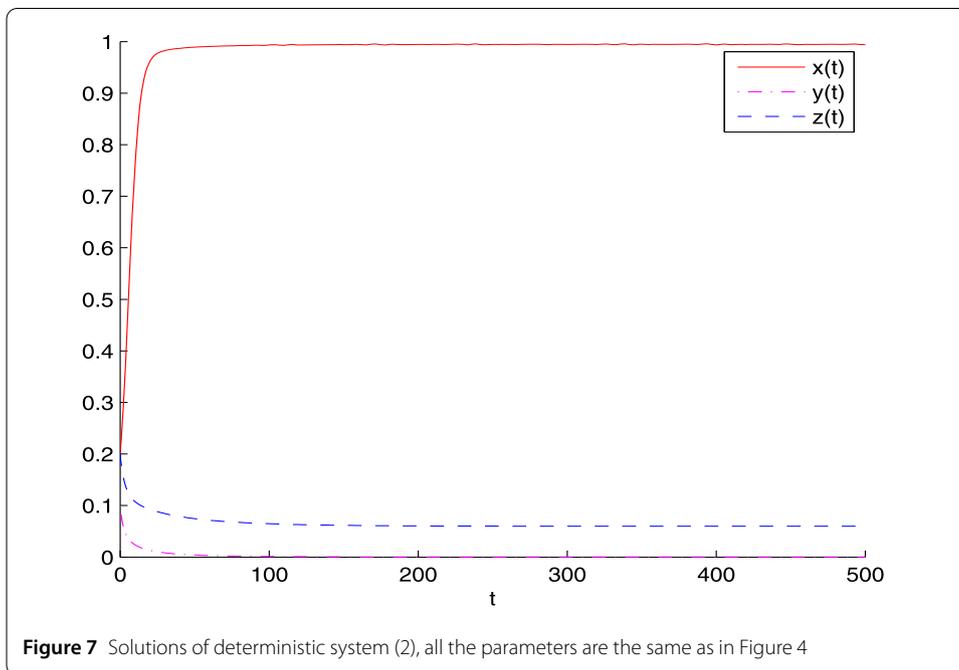
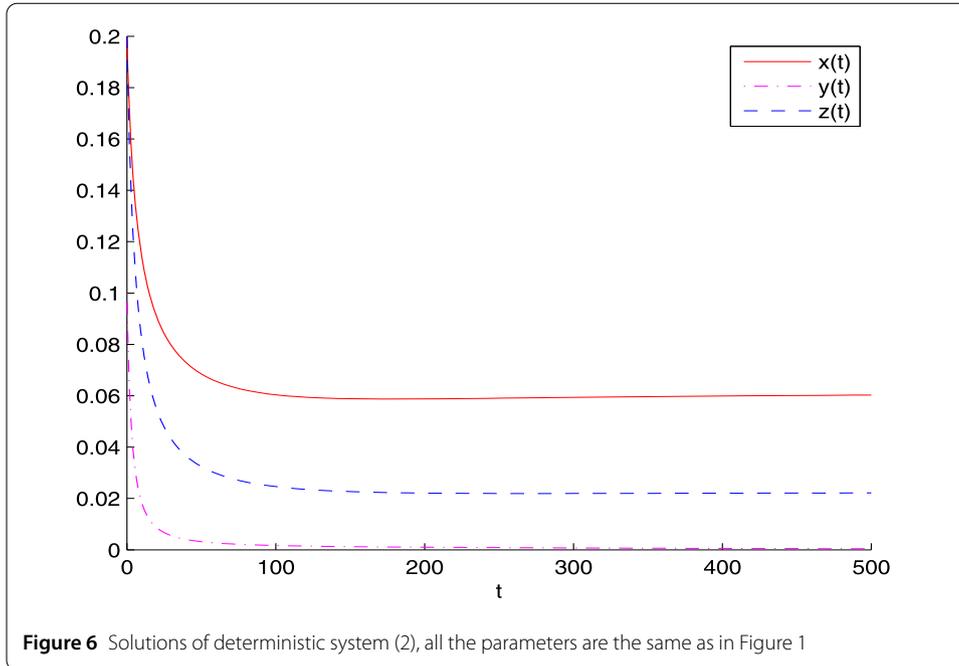
$$w_1 = 0.1573 > 0, \quad w_2 = -0.1585 < 0, \quad w_3 = -0.1685 < 0.$$



On the basis of (iv) in Theorem 2, the prey population will be stable in the mean, the whole predator population will die out, which is consistent with Figure 4.

- (5) In Figure 5, let  $a_1 = 0.75, a_2 = 0.7, a_3 = 0.3, c_2 = 0.3, c_3 = 0.2, k_2 = 0.6, \beta = 0.4, \sigma_3 = 0.7, \gamma_1(u) = 0.4, \gamma_2(u) = \gamma_3(u) = 0.2$ , the other parameters are the same as in (23), then

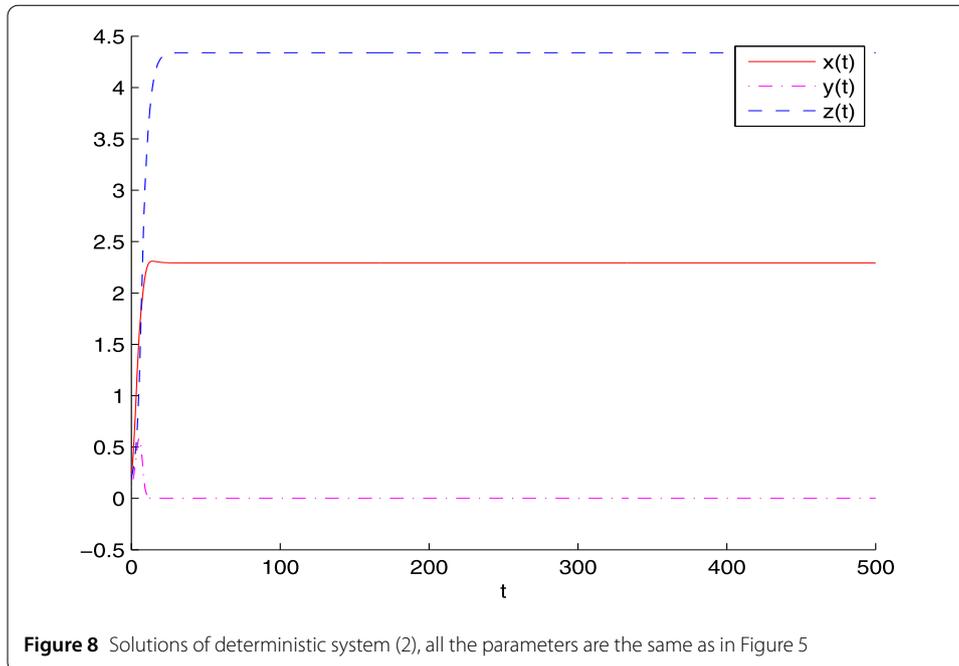
$$w_1 = 0.5615 > 0, \quad w_2 = 0.5573 > 0, \quad w_3 = 0.0373 > 0,$$



and  $\frac{(w_2+w_3)k_2c_1}{(c_2+c_3+k_2\beta)k_1} = 0.3857$ ,  $w_1 > \frac{(w_2+w_3)k_2c_1}{(c_2+c_3+k_2\beta)k_1}$ . By virtue of (v) in Theorem 2, all the populations will be strongly persistent in the mean.

In the following, we discuss the solution of deterministic system (2), assume that all the parameters are the same as in Figures 1, 4, 5. Figure 6 shows that the susceptible predator population will be extinct, but the prey and infected predator populations will be stable, which is different from Figure 1.

Figure 7 illustrates that the prey population and the infected predator population will be stable, and the susceptible predator population will go to extinction. Compare Figure 7



with Figure 4, the behavior of solution for system (2) is distinct from that for stochastic system (3). Figure 8 confirms it again, we give no more explanations.

These results show that, for a stochastic system, not only the growth rate but also the Lévy noise and Gauss white noise play an important role in the persistence of population. Sometimes, the stochastic perturbation of environment can make species extinct.

## 6 Conclusions

In this paper, we discuss an LG–Holling type II diseased predator ecosystem with Lévy noise and Gauss white noise. We show that the system admits a unique global positive solution; we also investigate persistence in the mean and extinction of all the populations, uniformly finite  $p$ th moment with  $p > 0$  and stochastic ultimate boundedness. The threshold conditions of extinction or persistence in the mean for prey and predator populations illustrate that  $w_i$  is the key value, which means that  $a_i$  is advantageous to the population, white noise and Lévy noise will go against the persistence of population. From the numerical simulation, we find that Lévy noise and Gauss white noise have an important effect on persistence for a stochastic system; occasionally, it may be a determinant factor. Thus, we must consider random changes of environment in the mathematical model.

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### Competing interests

The author declares that no competing interests exist.

### Authors' contributions

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## References

1. Leslie, P.H.: Some further notes on the use of matrices in population mathematics. *Biometrika* **35**, 213–245 (1948)
2. Leslie, P.H.: A stochastic model for studying the properties of certain biological systems numerical methods. *Biometrika* **45**, 16–31 (1958)
3. Pielou, E.C.: *Mathematical Ecology*. Wiley, New York (1977)
4. Korobeinikov, A.: A Lyapunov function for Leslie–Gower predator–prey models. *Appl. Math. Lett.* **14**, 697–699 (2001)
5. Zhou, X., Cui, J., Shi, X., Song, X.: A modified Leslie–Gower predator–prey model with prey infection. *J. Appl. Math. Comput.* **33**, 471–487 (2010)
6. Liu, M., Wang, K.: Dynamics of a Leslie–Gower Holling-type II predator–prey system with Lévy jumps. *Nonlinear Anal.* **85**, 204–213 (2013)
7. Flores, J.D., Gonzalez-Olivares, E.: A modified Leslie–Gower predator–prey model with ratio-dependent functional response and alternative food for the predator. *Math. Methods Appl. Sci.* **40**, 2313–2328 (2017)
8. Xu, Y., Liu, M., Yang, Y.: Analysis of a stochastic two-predators one-prey system with modified Leslie–Gower and Holling-type II schemes. *J. Math. Anal. Appl.* **7**, 713–727 (2017)
9. Sarwardi, S., Haque, M., Venturino, E.: Global stability and persistence in LG–Holling type II diseased predator ecosystems. *J. Biol. Phys.* **37**, 91–106 (2011)
10. Bao, J., Mao, X., Yin, G., Yuan, C.: Competitive Lotka–Volterra population dynamics with jumps. *Nonlinear Anal.* **74**, 6601–6616 (2011)
11. Bao, J., Yuan, C.: Stochastic population dynamics driven by Lévy noise. *J. Math. Anal. Appl.* **391**, 363–375 (2012)
12. Liu, M., Deng, M., Do, B.: Analysis of a stochastic logistic model with diffusion. *Appl. Math. Comput.* **266**, 169–182 (2015)
13. Liu, Q., Chen, Q.: Asymptotic behavior of a stochastic non-autonomous predator–prey system with jumps. *Appl. Math. Comput.* **271**, 418–428 (2015)
14. Liu, M., Wang, K.: Stochastic Lotka–Volterra systems with Levy noise. *J. Math. Anal. Appl.* **410**, 750–763 (2014)
15. Zhao, Y., Yuan, S.: Optimal harvesting policy of a stochastic two-species competitive model with Levy noise in a polluted environment. *Physica A* **477**, 20–33 (2017)
16. Liu, L., Meng, X., Zhang, T.: Optimal control strategy for an impulsive stochastic competition system with time delays and jumps. *Physica A* **477**, 99–113 (2017)
17. Liu, Q., Jiang, D., Shi, N., Hayat, T., Alsaedi, A.: Stochastic mutualism model with Levy jumps. *Commun. Nonlinear Sci. Numer. Simul.* **43**, 78–90 (2017)
18. Peng, S., Zhu, X.: Necessary and sufficient condition for comparison theorem of 1-dimensional stochastic differential equations. *Stoch. Process. Appl.* **116**, 370–380 (2006)
19. Kunita, H.: Itô's stochastic calculus: its surprising power for applications. *Stoch. Process. Appl.* **120**, 622–652 (2010)
20. Protter, P., Talay, D.: The Euler scheme for Lévy driven stochastic differential equations. *Ann. Probab.* **25**, 393–423 (1997)

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